◆ □ ▶ ◆ □ ▶ ▲ □ ▶ ▲ □ ▶ ● ○ ○ ○ ○

Reconstructing stellar DEM and metallicity using high-resolution X-ray Spectra

Victoria Liublinska Joint work with CHASC: X.-L. Meng, V. Kashyap, D. van Dyk

Harvard University

June 15, 2010

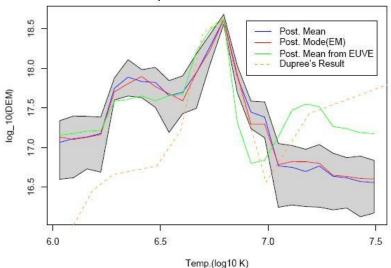
The model

Let $Y_i^{obs} \sim Pois(\xi_i)$, where ξ_i is photon intensity in wavelength channel *i*, *i* = 1...*I* and

$$\begin{split} \boldsymbol{\xi} &= \boldsymbol{\xi}^{source} + \boldsymbol{\xi}^{bkg} = MD\left(\boldsymbol{\lambda}^{C} + \sum_{k} \boldsymbol{\lambda}_{k}^{L}\right) + \boldsymbol{\xi}^{bkg} \\ &= MD\left(G^{C} + \sum_{k} \boldsymbol{\gamma}_{k} G_{k}^{L}\right)\boldsymbol{\mu} + \boldsymbol{\xi}^{bkg}, \end{split}$$

Parameters are γ and μ .

Previously estimated Capella DEM



Capella DEM Reconstruction

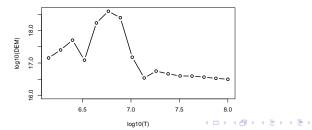
◆□▶ ◆□▶ ◆三▶ ◆三▶ ・三 ・ のへで

æ

log10(DEM) used for simulations

log10(DEM) used for simulations (16 nodes)





Other properties of current simulations

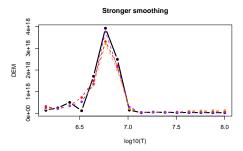
• Includes all censoring, adds informative prior to abundances

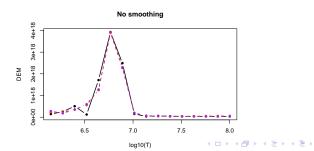
• 16 nodes and some smoothing (α_{ρ} =20): Depending on the starting point EM converges in 200-500 iterations

• 16 nodes and some smoothing (α_{ρ} =2): EM converges in 800-1100 iterations

æ

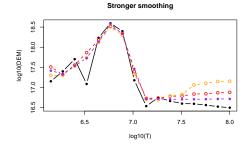
Estimated DEM

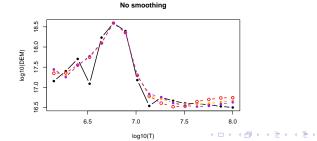




æ

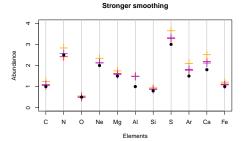
Estimated log10(DEM)



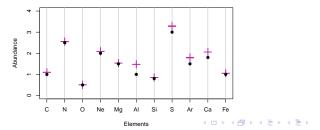


∃ 990

Estimated abundance

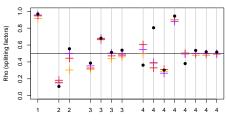


No smoothing



€ 990

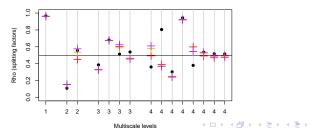
Estimated rho



Stronger smoothing

Multiscale levels



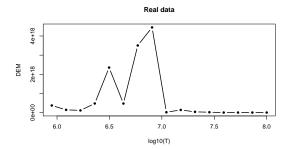


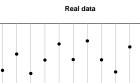
Abundance 0.0 0.4 0.8

С

N O Ne Ma

Estimated rho







Si S Ar

Ca Fe

Ni

▲□▶ ▲□▶ ▲ □▶ ▲ □▶ ▲ □ ● ● ● ●

◆ □ ▶ ◆ □ ▶ ▲ □ ▶ ▲ □ ▶ ● ○ ○ ○ ○

Gibbs in a nutshell

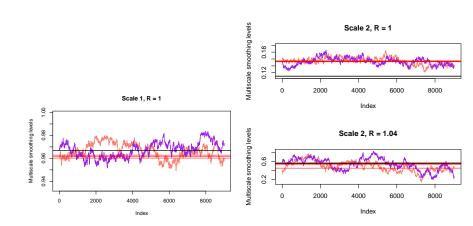
- Initiating parameters γ, μ
- 5 steps of sampling different missing data
- Updating abundance $\gamma^{(t+1)} | \mu^{(t)}, \gamma^{(t)}, Y_{mis}$
- 3 more steps of sampling missing data
- Updating (scaled) DEM $\tilde{\mu}^{(t+1)}|Y_{mis}$, where $\tilde{\mu} = f(\gamma)\mu$
- Rescale DEM $\mu^{(t+1)} = \tilde{\mu}^{(t+1)} / f(\gamma^{(t+1)})$

◆ □ ▶ ◆ □ ▶ ▲ □ ▶ ▲ □ ▶ ● ○ ○ ○ ○

Gibbs in a nutshell

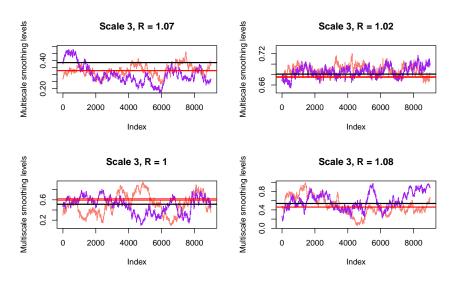
- Initiating parameters γ, μ
- 5 steps of sampling different missing data
- Updating abundance $\gamma^{(t+1)}|\mu^{(t)},\gamma^{(t)},Y_{mis}$
- 3 more steps of sampling missing data
- Updating (scaled) DEM $\tilde{\mu}^{(t+1)}|Y_{mis}$, where $\tilde{\mu} = f(\gamma)\mu$
- Rescale DEM $\mu^{(t+1)} = \tilde{\mu}^{(t+1)} / f(\gamma^{(t+1)})$
- Properties of a current simulation: minimal smoothing (α_ρ=2), 16 nodes for *log*10(*T*), 10000 simulations including 1000 of burn-in

Rho trace



◆□▶ ◆□▶ ◆三▶ ◆三▶ ・三三 - のへで

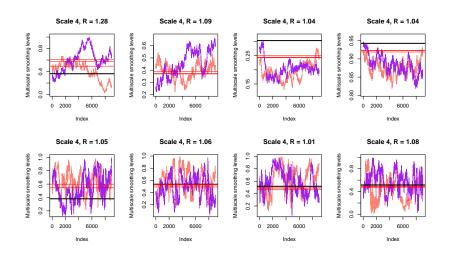
Rho trace



◆□▶ ◆□▶ ◆三▶ ◆三▶ ・三 ・ のへで

Alternative sampling of the DEM

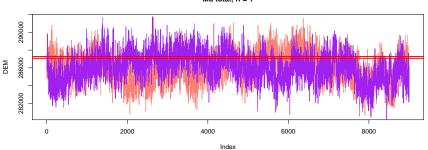
Rho trace



◆□▶ ◆□▶ ◆ □ ▶ ◆ □ ▶ ● □ ● ● ● ●

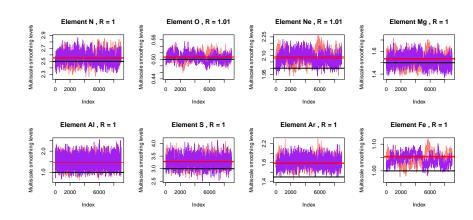
イロト イポト イモト イモ

Mu total trace



Mu total, R = 1

Abundance trace



◆□▶ ◆□▶ ◆臣▶ ◆臣▶ ─臣 ─のへで

◆□▶ ◆□▶ ◆ □▶ ◆ □▶ ○ □ ○ ○ ○ ○

Posterior distribution of $\mu | \gamma, Y^{obs}$

$$p(\boldsymbol{\mu}|\boldsymbol{\gamma}, Y^{obs}) \propto p(\boldsymbol{\mu}) \prod_{i} p(Y_{i}^{obs}|\boldsymbol{\mu}, \boldsymbol{\gamma}),$$

where

$$\begin{split} Y_i^{obs} | \boldsymbol{\mu}, \boldsymbol{\gamma} &\sim Pois\left(\sum_{j=1}^J M_{i,j} d_j \left(\sum_{t=1}^T G_{j,t}^C \boldsymbol{\mu}_t + \sum_{t=1}^T \sum_k \gamma_k G_{k,j,t}^L \boldsymbol{\mu}_t\right) + \boldsymbol{\xi}_i^{bkg}\right) \\ &= Pois\left(\sum_{t=1}^T h_{i,t} \boldsymbol{\mu}_t + \boldsymbol{\xi}_i^{bkg}\right), \end{split}$$

Posterior distribution of $\mu | \gamma, Y^{obs}$

$$p(\boldsymbol{\mu}|\boldsymbol{\gamma}, Y^{obs}) \propto p(\boldsymbol{\mu}) \prod_{i} p(Y_{i}^{obs}|\boldsymbol{\mu}, \boldsymbol{\gamma}),$$

where

$$\begin{split} Y_i^{obs} | \boldsymbol{\mu}, \boldsymbol{\gamma} &\sim Pois\left(\sum_{j=1}^J M_{i,j} d_j \left(\sum_{t=1}^T G_{j,t}^C \boldsymbol{\mu}_t + \sum_{t=1}^T \sum_k \gamma_k G_{k,j,t}^L \boldsymbol{\mu}_t\right) + \boldsymbol{\xi}_i^{bkg}\right) \\ &= Pois\left(\sum_{t=1}^T h_{i,t} \boldsymbol{\mu}_t + \boldsymbol{\xi}_i^{bkg}\right), \end{split}$$

with $\mu_1 = \mu_0 \prod_{r=1}^R \rho_{r,0}$, $\mu_T = \mu_0 \prod_{r=1}^R (1 - \rho_{r,2^r-1})$ and each μ_t is a product of μ_0 and *R* splitting factors, either $\rho_{r,n}$ or $(1 - \rho_{r,n})$.

◆□▶ ◆圖▶ ◆臣▶ ◆臣▶ 臣 - のへぐ

Posterior distribution of $\mu | \gamma, Y^{obs}$

$$p(\boldsymbol{\mu}|\boldsymbol{\gamma}, Y^{obs}) \propto p(\boldsymbol{\mu}) \prod_{i} p(Y_{i}^{obs}|\boldsymbol{\mu}, \boldsymbol{\gamma}),$$

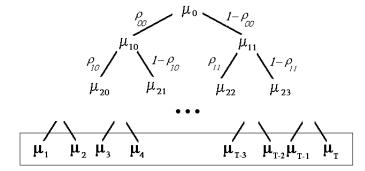
where

$$\begin{split} Y_i^{obs} | \boldsymbol{\mu}, \boldsymbol{\gamma} \sim Pois\left(\sum_{j=1}^J M_{i,j} d_j \left(\sum_{t=1}^T G_{j,t}^C \boldsymbol{\mu}_t + \sum_{t=1}^T \sum_k \gamma_k G_{k,j,t}^L \boldsymbol{\mu}_t\right) + \boldsymbol{\xi}_i^{bkg}\right) \\ = Pois\left(\sum_{t=1}^T h_{i,t} \boldsymbol{\mu}_t + \boldsymbol{\xi}_i^{bkg}\right), \end{split}$$

with $\mu_1 = \mu_0 \prod_{r=1}^R \rho_{r,0}$, $\mu_T = \mu_0 \prod_{r=1}^R (1 - \rho_{r,2^r-1})$ and each μ_t is a product of μ_0 and *R* splitting factors, either $\rho_{r,n}$ or $(1 - \rho_{r,n})$. The following example clarifies the representation (R=4)

$$\mu_{12} = \mu_0 q_{12,R} = \mu_0 (1 - \rho_{0,0}) q_{12,R-1} = \mu_0 (1 - \rho_{0,0}) \rho_{1,1} q_{12,R-2}$$
$$= \dots = \mu_0 (1 - \rho_{0,0}) \rho_{1,1} (1 - \rho_{2,3}) (1 - \rho_{2,6})$$

Parametrization of the multiscale smoothing



◆□▶ ◆□▶ ◆臣▶ ◆臣▶ ─臣 ─のへ⊙

Posterior distribution of $(\mu_0, \rho) | \gamma, Y^{obs}$

- $\mu_0 \sim Gamma(\alpha_{\mu})/\beta_{\mu}$, where flat prior would correspond to $\alpha_{\mu} = 1$ and $\beta_{\mu} = 0$
- $\rho_{r,n} \sim Beta(\alpha_{\rho}, \alpha_{\rho}), r = 0, ..., R 1 \text{ and } n = 0, ..., 2^r 1$
- Instead of working with Y^{obs} we can also use Y (counts free from the background contamination)

Posterior distribution of $(\mu_0, \rho) | \gamma, Y^{obs}$

- $\mu_0 \sim Gamma(\alpha_{\mu})/\beta_{\mu}$, where flat prior would correspond to $\alpha_{\mu} = 1$ and $\beta_{\mu} = 0$
- $\rho_{r,n} \sim Beta(\alpha_{\rho}, \alpha_{\rho}), r = 0, ..., R 1 \text{ and } n = 0, ..., 2^r 1$
- Instead of working with Y^{obs} we can also use Y (counts free from the background contamination)

$$p(\mu_0, \boldsymbol{\rho} | \boldsymbol{\gamma}, \boldsymbol{Y}) \propto p(\mu_0) \prod_{r,n} p(\rho_{r,n}) \prod_i p(Y_i | \boldsymbol{\rho}, \mu_0, \boldsymbol{\gamma}).$$

Posterior distribution of $(\mu_0, \rho) | \gamma, Y^{obs}$

- $\mu_0 \sim Gamma(\alpha_{\mu})/\beta_{\mu}$, where flat prior would correspond to $\alpha_{\mu} = 1$ and $\beta_{\mu} = 0$
- $\rho_{r,n} \sim Beta(\alpha_{\rho}, \alpha_{\rho}), r = 0, ..., R 1 \text{ and } n = 0, ..., 2^r 1$
- Instead of working with Y^{obs} we can also use Y (counts free from the background contamination)

$$p(\mu_0, \boldsymbol{\rho} | \boldsymbol{\gamma}, \boldsymbol{Y}) \propto p(\mu_0) \prod_{r,n} p(\rho_{r,n}) \prod_i p(Y_i | \boldsymbol{\rho}, \mu_0, \boldsymbol{\gamma}).$$

As it was noted previously

$$Y_i|\mu_0, \boldsymbol{\rho}, \boldsymbol{\gamma} \sim Pois\left(\sum_{t=1}^T h_{i,t}\mu_t\right) = Pois\left(\mu_0 \sum_{t=1}^T h_{i,t}q_{t,R}\right) = Pois\left(s_i\mu_0\right)$$

・ロト・日本・日本・日本・日本・日本

Posterior distribution of $(\mu_0, \rho) | \gamma, Y^{obs}$

- $\mu_0 \sim Gamma(\alpha_\mu)/\beta_\mu$, where flat prior would correspond to $\alpha_\mu = 1$ and $\beta_\mu = 0$
- $\rho_{r,n} \sim Beta(\alpha_{\rho}, \alpha_{\rho}), r = 0, ..., R 1 \text{ and } n = 0, ..., 2^r 1$
- Instead of working with Y^{obs} we can also use Y (counts free from the background contamination)

$$p(\mu_0, \boldsymbol{\rho} | \boldsymbol{\gamma}, \boldsymbol{Y}) \propto p(\mu_0) \prod_{r,n} p(\rho_{r,n}) \prod_i p(Y_i | \boldsymbol{\rho}, \mu_0, \boldsymbol{\gamma}).$$

As it was noted previously

$$Y_i|\mu_0, \boldsymbol{\rho}, \boldsymbol{\gamma} \sim Pois\left(\sum_{t=1}^T h_{i,t}\mu_t\right) = Pois\left(\mu_0 \sum_{t=1}^T h_{i,t}q_{t,R}\right) = Pois\left(s_i\mu_0\right)$$

Therefore, μ_0 can be updated in closed form without any missing data imputation

$$\mu_0 | \boldsymbol{\rho}, \boldsymbol{Y}, \boldsymbol{\gamma} \sim Gamma\left(\alpha_{\mu} + \sum_i Y_i\right) / \left(\beta_{\mu} + \sum_i s_i\right)$$

Updating splitting factors $t_{\alpha}(\alpha, \alpha)$ and

Since $\rho_{r,n} \sim Beta(\alpha_{\rho}, \alpha_{\rho})$ and

$$p(\rho_{0,0}|\boldsymbol{\rho}_{-},\mu_{0},\boldsymbol{\gamma},\boldsymbol{Y}) \propto p(\rho_{0,0}) \prod_{i} p(Y_{i}|\boldsymbol{\rho},\mu_{0},\boldsymbol{\gamma}).$$

The distribution of background-free counts can be represented as

$$Y_{i}|\mu_{0},\boldsymbol{\rho},\boldsymbol{\gamma} = Pois\left(\sum_{t=1}^{T/2} h_{i,t}\mu_{0}\rho_{0,0}q_{t,R-1} + \sum_{t=T/2+1}^{T} h_{i,t}\mu_{0}(1-\rho_{0,0})q_{t,R-1}\right)$$
$$= Pois\left(s_{i,1}\rho_{0,0} + s_{i,2}(1-\rho_{0,0})\right),$$

▲□▶ ▲□▶ ▲三▶ ▲三▶ 三三 のへの

Updating splitting factors ta(q, q) and

Since $\rho_{r,n} \sim Beta(\alpha_{\rho}, \alpha_{\rho})$ and

$$p(\rho_{0,0}|\boldsymbol{\rho}_{-},\mu_0,\boldsymbol{\gamma},\boldsymbol{Y}) \propto p(\rho_{0,0}) \prod_i p(Y_i|\boldsymbol{\rho},\mu_0,\boldsymbol{\gamma}).$$

The distribution of background-free counts can be represented as

$$Y_{i}|\mu_{0},\boldsymbol{\rho},\boldsymbol{\gamma} = Pois\left(\sum_{t=1}^{T/2} h_{i,t}\mu_{0}\rho_{0,0}q_{t,R-1} + \sum_{t=T/2+1}^{T} h_{i,t}\mu_{0}(1-\rho_{0,0})q_{t,R-1}\right)$$
$$= Pois\left(s_{i,1}\rho_{0,0} + s_{i,2}(1-\rho_{0,0})\right),$$

then the conditional distribution of $\rho_{0,0}|\rho_-,\mu_0,Y,\gamma$ has the following form

$$\rho_{0,0}^{\alpha_{\rho}-1}(1-\rho_{0,0})^{\alpha_{\rho}-1}\prod_{i}(s_{i,1}\rho_{0,0}+s_{i,2}(1-\rho_{0,0}))^{Y_{i}}e^{-(s_{i,1}\rho_{0,0}+s_{i,2}(1-\rho_{0,0}))}$$

Updating splitting factors

Possible idea: update splitting factors by scale level. The number of simultaneously updated factors would be 1, 2, 4, 8, 16,etc.

Updating splitting factors

Possible idea: update splitting factors by scale level. The number of simultaneously updated factors would be 1, 2, 4, 8, 16,etc. For example let

$$\begin{split} \eta_{i} &= s_{i,2,1}\rho_{1,0} + s_{i,2,2}(1-\rho_{1,0}) + s_{i,2,3}\rho_{1,1} + s_{i,2,4}(1-\rho_{1,1}), \\ \text{then } p(\rho_{1,0},\rho_{1,1}|\boldsymbol{\rho}_{-},\mu_{0},\boldsymbol{\gamma},\boldsymbol{Y}) \propto \\ \rho_{1,0}^{\alpha_{\rho}-1}(1-\rho_{1,0})^{\alpha_{\rho}-1}\rho_{1,1}^{\alpha_{\rho}-1}(1-\rho_{1,1})^{\alpha_{\rho}-1}\prod_{i}\eta_{i}^{Y_{i}}e^{-\eta_{i}} \end{split}$$

Updating splitting factors

Possible idea: update splitting factors by scale level. The number of simultaneously updated factors would be 1, 2, 4, 8, 16,etc. For example let

$$\begin{split} &\eta_i = s_{i,2,1}\rho_{1,0} + s_{i,2,2}(1-\rho_{1,0}) + s_{i,2,3}\rho_{1,1} + s_{i,2,4}(1-\rho_{1,1}), \\ & \text{then } p(\rho_{1,0},\rho_{1,1}|\boldsymbol{\rho}_{-},\mu_0,\boldsymbol{\gamma},\boldsymbol{Y}) \propto \\ &\rho_{1,0}^{\alpha_\rho-1}(1-\rho_{1,0})^{\alpha_\rho-1}\rho_{1,1}^{\alpha_\rho-1}(1-\rho_{1,1})^{\alpha_\rho-1}\prod_i \eta_i^{Y_i}e^{-\eta_i} \end{split}$$

Implementation questions:

- Which proposal is the best? Truncated MVN (or MVt), others? OR simply evaluate the posterior and sample from the grid (since each ρ_{r,n} < 1)?
- Other updating scheme? One-by-one, in pairs etc.
- Update μ directly? (the posterior is not that simple since each prior on $\rho_{r,n}$ contains all μ_1, \dots, μ_T)
- Alternate this scheme with full augmentation?